## Solution for HW3, MATH3805

1. Exercise 5.7
(a) (i) first-order;
(ii) third-order;
(iii) first-order;
(iv) second-order.
(b) (i) $E(y)=\beta_{0}+\beta_{1} x$
(ii) $E(y)=\beta_{0}+\beta_{1} x+\beta_{2} x^{2}+\beta_{3} x^{3}$
(iii) $E(y)=\beta_{0}+\beta_{1} x$
(iv) $E(y)=\beta_{0}+\beta_{1} x+\beta_{2} x^{2}$
(c) (i) $\beta_{1}>0$
(ii) $\beta_{3}>0$
(iii) $\beta_{1}<0$
(iv) $\beta_{2}<0$
2. Exercise 5.17
(a) $E(y)=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{4}+\beta_{4} x_{1} x_{2}+\beta_{5} x_{1} x_{4}+\beta_{6} x_{2} x_{4}+\beta_{7} x_{1}^{2}+\beta_{8} x_{2}^{2}+\beta_{9} x_{4}^{2}$
(b) $H_{0}: \beta_{1}=\beta_{2}=\ldots=\beta_{9}=0$ v.s. $H_{1}$ : at least one $\beta_{i} \neq 0$

F-test statistic is 613.27 (SAS). P -value $<0.0001<\alpha$.
so reject $H_{0}$. we have enough evidence to conclude that the overall model is statistically useful for predicting $y$.
(c) compare a reduced model:
$E(y)=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{4}+\beta_{4} x_{1} x_{4}+\beta_{5} x_{1} x_{4}+\beta_{6} x_{2} x_{4}$
$H_{0}: \beta_{7}=\beta_{8}=\beta_{9}=0$ v.s. $H_{1}:$ at least one $\beta_{i} \neq 0, i=7,8,9$
$F=\frac{\left(\frac{\left(\text { SSE }_{r}-\text { SSE }_{c}\right.}{g}\right)}{\left(\frac{\operatorname{SSE}_{c}}{n-(k+1)}\right)}=\frac{\left(\frac{3314.94919-146.12634}{3}\right)}{\left(\frac{1466.12633}{25-10}\right)}=108.4275$
so reject $H_{0}$. we have enough evidence to conclude that the curvilinear terms is statistically useful for predicting $y$.

| Analysis of Variance |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Source | DF | Sum of <br> Squares | Mean <br> Square | EValue | Pr 2 |
| Model | 9 | 53769 | 5974.36818 | 613.27 | $<.0001$ |
| Error | 15 | 146.12634 | 9.74176 |  |  |
| Corrected Total | 24 | 53915 |  |  |  |


| Root MSE | 3.12118 | R-Square | 0.9973 |
| :--- | ---: | :--- | :--- |
| Dependent Mean | 88.32000 | Adj R-Sq | 0.9957 |
| Coeff Var | 3.53394 |  |  |


| Parameter Estimates |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | :---: |
| Variable | DF | Parameter <br> Estimate | Standard <br> Error | t Value | Pr > \|t| |  |
| Intercept | 1 | 10283 | 2335.56968 | 4.40 | 0.0005 |  |
| REDSHIFT | 1 | 276.75689 | 147.03327 | 1.88 | 0.0793 |  |
| LINEFLUX | 1 | 3325.18799 | 476.53498 | 6.98 | $<.0001$ |  |
| AB1450 | 1 | 1301.30580 | 115.21604 | 11.29 | $<.0001$ |  |
| RL | 1 | 41.97757 | 14.32647 | 2.93 | 0.0103 |  |
| RA | 1 | 15.98078 | 4.23728 | 3.77 | 0.0018 |  |
| LA | 1 | 207.37869 | 11.82790 | 17.53 | $<.0001$ |  |
| REDSHIFTSQ | 1 | 0.98954 | 4.30238 | 0.23 | 0.8212 |  |
| LINEFLUXSQ | 1 | 266.55556 | 24.84031 | 10.73 | $<.0001$ |  |
| AB1450SQ | 1 | 40.21654 | 2.42805 | 16.56 | $<.0001$ |  |

Figure 1: SAS output without intercept for Exercise 5.17

| Analysis of Variance |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Source | DF | Sum of <br> Squares | Mean <br> Square | F Value | Pr $>$ F |$|$| Model | 6 | 50600 | 8433.41514 |
| :--- | ---: | ---: | ---: |
| Error | 18 | 3314.94919 | 184.16384 |
| Corrected Total | 24 | 53915 |  |


| Root MSE | 13.57070 | R-Square | 0.9385 |
| :--- | :--- | :--- | :--- |
| Dependent Mean | 88.32000 | Adj R-Sq | 0.9180 |
| Coeff Var | 15.36537 |  |  |


| Parameter Estimates |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | :---: |
| Variable | DF | Parameter <br> Estimate | Standard <br> Error | t Value | Pr > \|t| |  |
| Intercept | 1 | -5590.24207 | 4472.44114 | -1.25 | 0.2273 |  |
| REDSHIFT | 1 | -826.71928 | 468.64057 | -1.76 | 0.0947 |  |
| LINEFLUX | 1 | -216.24176 | 355.58876 | -0.61 | 0.5507 |  |
| AB1450 | 1 | 625.47619 | 223.35634 | 2.80 | 0.0118 |  |
| RL | 1 | -83.23183 | 41.00448 | -2.03 | 0.0574 |  |
| RA | 1 | -15.71527 | 13.16828 | -1.19 | 0.2482 |  |
| LA | 1 | 35.57030 | 17.14111 | 2.08 | 0.0526 |  |

Figure 2: SAS output without intercept for Exercise 5.17
3. Exercise 5.20
(a) $\bar{x}=33, s_{x}=2.16025$.

The coding system equation: $u=(x-33) / 2.16025$
(b) $x=[30,31,32,33,34,35,36]$
$\rightarrow u=[-1.38873,-0.92582,-0.46291,0,0.46291,0.92582,1.38873]$
(c) $\operatorname{cor}\left(x, x^{2}\right)=\frac{\operatorname{cov}\left(x, x^{2}\right)}{\sqrt{\operatorname{var}(x) \operatorname{var}\left(x^{2}\right)}}=0.99966$
(d) $\operatorname{cor}\left(u, u^{2}\right)=0$
(e) $E(y)=37.57143-0.46291 u-5.3333 u^{2}$

| Analysis of Variance |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Source | DF | Sum of <br> Squares | Mean <br> Square | F Value | $\operatorname{Pr}>$ F |
| Model | 2 | 111.00000 | 55.50000 | 44.40 | 0.0019 |
| Error | 4 | 5.00000 | 1.25000 |  |  |
| Corrected Total | 6 | 116.00000 |  |  |  |


| Root MSE | 1.11803 | R-Square | 0.9569 |
| :--- | ---: | :--- | :--- |
| Dependent Mean | 33.00000 | Adj R-Sq | 0.9353 |
| Coeff Var | 3.38798 |  |  |


| Parameter Estimates |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | :---: |
| Variable | DF | Parameter <br> Estimate | Standard <br> Error | t Value | Pr > \|t| |  |
| Intercept | 1 | 37.57143 | 0.64550 | 58.21 | $<.0001$ |  |
| $\mathbf{u}$ | 1 | -0.46291 | 0.45644 | -1.01 | 0.3679 |  |
| usq | 1 | -5.33333 | 0.56928 | -9.37 | 0.0007 |  |

Figure 3: SAS output without intercept for Exercise 5.20

## 4. Exercise 5.24

(a) $E(y)=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}$
where $x_{1}=\left\{\begin{array}{ll}1, & \text { male ; } \\ 0, & \text { female. }\end{array}\right.$ and $x_{2}= \begin{cases}1, & \text { expert testimont }=\text { yes } \\ 0, & \text { otherwise }\end{cases}$
Interpretation of model parameters:
$\beta_{0}=u_{11}$ mean at gender $=$ female and expert testimony $=$ yes
$\beta_{1}=u_{2 j}-u_{1 j}$, for any level of expert testimony
$\beta_{2}=u_{i 2}-u_{i 1}$, for any level of gender
(b) $E(y)=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{1} x_{2}$

Interpretation of model parameters:
$\beta_{0}=u_{11}$ mean at gender -female and expert testimony $=$ yes
$\beta_{1}=u_{2 j}-u_{1 j}$, for any level of expert testimony
$\beta_{2}=u_{i 2}-u_{i 1}$, for any level of gender
$\beta_{3}=u_{22}-u_{12}-u_{21}+u_{11}$,
(c) Model in part (b).
when $x_{1}=0, x_{2}=1, E(y)=\beta_{0}+\beta_{2}$
when $x_{1}=1, x_{2}=1, E(y)=\beta_{0}+\beta_{1}+\beta_{2}+\beta_{3}$
when $x_{1}=0, x_{2}=0, E(y)=\beta_{0}$
when $x_{1}=1, x_{2}=0, E(y)=\beta_{0}+\beta_{1}$
and we assume that

$$
\begin{aligned}
& u_{12}>u_{22} \Rightarrow \beta_{0}+\beta_{2}>\beta_{0}+\beta_{1}+\beta_{2}+\beta_{3} \\
& u_{21}>u_{21} \Rightarrow \beta_{0}<\beta_{0}+\beta_{1}
\end{aligned}
$$

Thus,

$$
\begin{aligned}
& \beta_{1}>0 \\
& \beta_{1}+\beta_{3}<0
\end{aligned}
$$

5. Exercise 5.34
(a)

$$
\begin{aligned}
E(y)= & \beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{3}+\beta_{4} x_{4} \\
& +\beta_{5} x_{1} x_{2}+\beta_{6} x_{1} x_{3}+\beta_{7} x_{1} x_{4}+\beta_{8} x_{2} x_{3}+\beta_{9} x_{2} x_{4} \\
& +\beta_{10} x_{1}^{2}+\beta_{11} x_{2}^{2}
\end{aligned}
$$

where $x_{1}$ is cycle speed, $x_{2}$ is cycle pressure ratio, take traditional level as base level.
$x_{3}=\left\{\begin{array}{ll}1, & \text { at advanced level; } \\ 0, & \text { not at advanced level. }\end{array}\right.$ and $x_{4}= \begin{cases}1, & \text { at aeroderivative level; } \\ 0, & \text { not at aeroderivative level. }\end{cases}$
(b) for traditional level, $x_{3}=x_{4}=0$, model is

$$
E(y)=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{5} x_{1} x_{2}+\beta_{10} x_{1}^{2}+\beta_{11} x_{2}^{2}
$$

for advanced level, $x_{3}=1, x_{4}=0$, model is

$$
\begin{aligned}
E(y)= & \beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{3} \\
& +\beta_{5} x_{1} x_{2}+\beta_{6} x_{1} x_{3}+\beta_{8} x_{2} x_{3} \\
& +\beta_{10} x_{1}^{2}+\beta_{11} x_{2}^{2}
\end{aligned}
$$

for aeroderivative level, $x_{3}=0, x_{4}=1$, model is

$$
\begin{aligned}
E(y)= & \beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{4} x_{4} \\
& +\beta_{5} x_{1} x_{2}+\beta_{7} x_{1} x_{4}+\beta_{9} x_{2} x_{4} \\
& +\beta_{10} x_{1}^{2}+\beta_{11} x_{2}^{2}
\end{aligned}
$$

(c)

$$
\begin{aligned}
E(y)= & 14485+0.11816 x_{1}-370.44605 x_{2}-2662.13217 x_{3}-1763.69024 x_{4} \\
& +0.00054906 x_{1} x_{2}+0.17531 x_{1} x_{3}+0.03417 x_{1} x_{4} \\
& +96.17827 x_{2} x_{3}+78.05405 x_{2} x_{4}-2.05897 \mathrm{E}-7 x_{1}^{2}+4.68457 x_{2}^{2}
\end{aligned}
$$

(d) $H_{0}: \beta_{1}=\beta_{2}=\ldots=\beta_{11}$ v.s. $H_{1}$ : at least one is not. F-value is 44.22 , and p -value is less than 0.0001
we have enough evidence to conclude that the overall model is useful.
(e) $H_{0}: \beta_{3}=\beta_{4}=\beta_{6}=\beta_{7}=\beta_{8}=\beta_{9}=0$ v.s. $H_{1}$ : at least one is not. $F=\frac{\left(\frac{\left(\text { SSE }_{r}-\text { SSE }_{c}\right.}{g}\right)}{\left(\frac{\operatorname{SE} E_{c}}{n-(k+1)}\right)}=\frac{\left(\frac{19370350-17056349}{6}\right)}{\left(\frac{1756399}{67-12}\right)}=1.243623$
we do not have enough evidence to conclude that the surfaces are not identical.

| Analysis of Variance |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Source | DF | Sum of <br> Squares | Mean <br> Square | F Value | Pr > F |
| Model | 5 | 148526859 | 29705372 | 93.55 | $<.0001$ |
| Error | 61 | 19370350 | 317547 |  |  |
| Corrected Total | 66 | 167897208 |  |  |  |


| Root MSE | 563.51284 | R-Square | 0.8846 |
| :--- | ---: | :--- | :--- |
| Dependent Mean | 11066 | Adj R-Sq | 0.8752 |
| Coeff Var | 5.09209 |  |  |


| Parameter Estimates |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | :---: |
| Variable | DF | Parameter <br> Estimate | Standard <br> Error | t Value | $\operatorname{Pr}>\|\mathrm{t}\|$ |  |
| Intercept | 1 | 15583 | 1142.85985 | 13.63 | $<.0001$ |  |
| X1 | 1 | 0.07823 | 0.11044 | 0.71 | 0.4814 |  |
| X2 | 1 | -523.13391 | 103.37571 | -5.06 | $<.0001$ |  |
| X1X2 | 1 | 0.00445 | 0.00558 | 0.80 | 0.4282 |  |
| X12 | 1 | $-1.80598 \mathrm{E}-7$ | 0.00000197 | -0.09 | 0.9272 |  |
| X22 | 1 | 8.84007 | 2.16320 | 4.09 | 0.0001 |  |

Figure 4: SAS output for Exercise 5.34

| Analysis of Variance |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Source | DF | Sum of <br> Squares | Mean <br> Square | F Value | Pr > F |
| Model | 11 | 150840860 | 13712805 | 44.22 | $<.0001$ |
| Error | 55 | 17056349 | 310115 |  |  |
| Corrected Total | 66 | 167897208 |  |  |  |


| Root MSE | 556.88009 | R-Square | 0.8984 |
| :--- | ---: | :--- | :--- |
| Dependent Mean | 11066 | Adj R-Sq | 0.8781 |
| Coeff Var | 5.03216 |  |  |


| Parameter Estimates |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | :---: |
| Variable | DF | Parameter <br> Estimate | Standard <br> Error | t Value | $\operatorname{Pr}>\|\mathbf{t}\|$ |  |
| Intercept | 1 | 14485 | 1897.87728 | 7.63 | $<.0001$ |  |
| X1 | 1 | 0.11816 | 0.14123 | 0.84 | 0.4064 |  |
| X2 | 1 | -370.44605 | 213.99588 | -1.73 | 0.0890 |  |
| X3 | 1 | -2662.13217 | 1458.64130 | -1.83 | 0.0734 |  |
| X4 | 1 | -1763.69024 | 3249.94564 | -0.54 | 0.5895 |  |
| X1X2 | 1 | 0.00054906 | 0.00727 | 0.08 | 0.9400 |  |
| X12 | 1 | $-2.05897 \mathrm{E}-7$ | 0.00000267 | -0.08 | 0.9388 |  |
| X22 | 1 | 4.68457 | 6.37404 | 0.73 | 0.4655 |  |
| X1X3 | 1 | 0.17531 | 0.09531 | 1.84 | 0.0713 |  |
| X2X3 | 1 | 96.14827 | 84.09666 | 1.14 | 0.2579 |  |
| X1X4 | 1 | 0.03417 | 0.07590 | 0.45 | 0.6543 |  |
| X2X4 | 1 | 78.05405 | 161.90057 | 0.48 | 0.6316 |  |

Figure 5: SAS output for Exercise 5.34

## 6. Exercise 5.35

(a) $E(y)=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{1} x_{2}$
(b) when $x_{2}=0, E(y)=\beta_{0}+\beta_{1} x_{1}$
when $x_{2}=1, E(y)=\beta_{0}+\beta_{2}+\left(\beta_{1}+\beta_{3}\right) x_{1}$
(c) when $x_{2}=1$, the change of $y$ for every one foot increse in elevation for moss specimens is $\beta_{1}+\beta_{3}$.
(d) $E(y)=2.38487+0.00181 x_{1}+3.20146 x_{2}-0.00133 x_{1} x_{2}$
$H_{0}: \beta_{1}=\beta_{2}=\beta_{3}=0$ v.s. $H_{1}$ : at least one is not.
F -value is 0.26 , and p -value is $0.8567>0.1$
we do not have enough evidence to conclude that the model is useful.
(e) $E(y)=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{1}^{2}+\beta_{3} x_{2}+\beta_{4} x_{1} x_{2}$

| Analysis of Variance |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Source | DF | Sum of <br> Squares | Mean <br> Square | F Value | Pr > F |
| Model | 3 | 20.23864 | 6.74621 | 0.26 | 0.8567 |
| Error | 66 | 1738.08227 | 26.33458 |  |  |
| Corrected Total | 69 | 1758.32091 |  |  |  |


| Root MSE | 5.13172 | R-Square | 0.0115 |
| :--- | ---: | :--- | :--- |
| Dependent Mean | 6.83994 | Adj R-Sq | -0.0334 |
| Coeff Var | 75.02582 |  |  |


| Parameter Estimates |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | :---: |
| Variable | DF | Parameter <br> Estimate | Standard <br> Error | $\mathbf{t}$ Value | $\operatorname{Pr}>\|\mathbf{t}\|$ |  |
| Intercept | 1 | 2.38487 | 5.39314 | 0.44 | 0.6598 |  |
| ELEVATION | 1 | 0.00181 | 0.00214 | 0.84 | 0.4014 |  |
| SLOPE | 1 | 3.20146 | 7.66988 | 0.42 | 0.6777 |  |
| ES | 1 | -0.00133 | 0.00303 | -0.44 | 0.6629 |  |

Figure 6: SAS output for Exercise 5.35
7. Exercise 5.37
(a) $E(y)=\beta_{0}++\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{3}$
(b) $E(y)=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{3}+\beta_{4} x_{1} x_{2}+\beta_{5} x_{1} x_{3}$
(c) For level AL, the slope of $E(y)$ is $\beta_{1}$

For level 3A, the slope of $E(y)$ is $\beta_{1}+\beta_{4}$
For level FE, the slope of $E(y)$ is $\beta_{1}+\beta_{5}$
(d) $H_{0}: \beta_{4}=\beta_{5}=0$ v.s. $H_{1}$ : at least one is not.
fit complete model and reduced model then use F test.
8. Exercise 5.45
(a) $\mathbb{E}(y)=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{3}$
where $x_{1}=\left\{\begin{array}{ll}1, & \text { boy } ; \\ 0, & \text { girl. }\end{array}, x_{2}=\left\{\begin{array}{ll}1, & \text { ygungest } ; \\ 0, & \text { not. }\end{array}, x_{3}= \begin{cases}1, & \text { middle } \\ 0, & \text { mnot. }\end{cases}\right.\right.$
(b) $\beta_{0}:$ mean at level: girl, oldest.
$\beta_{1}$ : mean difference between firl and boy.
$\beta_{2}$ : mean difference between youngest and oldest.
$\beta_{3}$ : mean difference between middle and oldest.
(c) $E(y)=\beta_{0}+\beta_{1} y_{1}+\beta_{2} x_{2}+\beta_{3} x_{3}+\beta_{4} x_{1} x_{2}+\beta_{5} x_{1} x_{3}$
(d) $\beta_{0}=0.21$
$\beta_{1}=0.16-0.21=-0.05$
$\beta_{2}=0.27-0.21=0.06$
$\beta_{3}=0.18-0.21=-0.03$
$\beta_{4}=0.33-0.27-0.16+0.21=0.11$
$\beta_{5}=0.33-0.18-0.16+0.21=0.20$
(e) $H_{0}: \beta_{4}=\beta_{5}=0$ v.s. $H_{1}$ : at least one is not.
9. Exercise 5.46
(a) $E(y)=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}$
(b) $\beta_{0}:$ mean at level: female, 1880.
$\beta_{1}$ : the slope for winning time.
$\beta_{2}$ : the mean difference of male and female.
(c) $E(y)=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x x_{2}+\beta_{3} x_{1} x_{2}$
for male: $E(y)=\left(\beta_{0} \not+\beta_{2}\right)+\left(\beta_{1}+\beta_{3}\right) x_{1}$
for female: $E(y)=\beta_{0}+\beta_{1} x_{1}$
(d) $E(y)=\beta_{0}+\beta_{1} x_{2}+\beta_{2} x_{3}+\beta_{3} x_{1} x_{3}+\beta_{4} x_{1}^{2}+\beta_{5} x_{3}^{2}$
$\begin{aligned} \text { (e) } E(y)= & \beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{3}+\beta_{3} x_{1} x_{3}+\beta_{4} x_{1}^{2}+\beta_{5} x_{3}^{2}+\beta_{6} x_{2} \\ \text { (f) } & \\ & \begin{aligned} & \\ & E(y)=\begin{array}{l}\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{3}+\beta_{3} x_{1} x_{3}+\beta_{4} x_{1}^{2}+\beta_{5} x_{3}^{2}+\beta_{6} x_{2} \\ \\ \end{array} \beta_{7} x_{1} x_{2}+\beta_{8} x_{3} x_{2}+\beta_{9} x_{1} x_{3} x_{2}+\beta_{10} x_{1}^{2} x_{2}+\beta_{11} x_{3}^{2} x_{2}\end{aligned}\end{aligned}$
(g)

$$
\begin{aligned}
E(y)= & \left(\beta_{0}+\beta_{1} x_{1}+\beta_{4} x_{1}^{2}\right)+\left(\beta_{2}+\beta_{3} x_{1}\right) x_{3}+\beta_{5} x_{3}^{2} \\
& \left.+\left(\beta_{6}+\beta_{7} x\right)+\beta_{10} x_{1}^{2}\right) x_{2}+\left(\beta_{8}+\beta_{9} x_{1}\right) x_{2} x_{3}+\beta_{11} x_{3}^{2} x_{2}
\end{aligned}
$$

The interaction of $x_{2}$ and $x_{3}$ should be deleted. So $\beta_{8}+\beta_{9} x_{1}=0, \beta_{11}=0$ and $\beta_{6}+\beta_{7} x_{1}+\beta_{10} x_{1}^{2} \neq 0$
(h) $\beta_{8}+\beta_{9} x_{1}=0, \beta_{11}=0$
(i) $\beta_{8}+\beta_{9} x_{1}=0, \beta_{11}=0$ and $\beta_{6}+\beta_{7} x_{1}+\beta_{10} x_{1}^{2}=0$

## IV. Exercise 5.47

(a) $E(y)=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{3}+\beta_{4} x_{4}$
where $x_{1}=\left\{\begin{array}{ll}1, & \mathrm{Pl} ; \\ 0, & \mathrm{P} 2\end{array}, x_{2}=\left\{\begin{array}{ll}1, & \mathrm{~L} 1 ; \\ 0, & \text { not. }\end{array} \quad, x_{3}=\left\{\begin{array}{ll}1, & \mathrm{~L} 2 ; \\ 0, & \text { not. }\end{array}, x_{4}=\left\{\begin{array}{cc}1, & \mathrm{~L} 4 \\ 0, & \text { not. }\end{array}\right.\right.\right.\right.$
(b) 8 parameters, $E(y)=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{3}+\beta_{4} x_{4}+\beta_{5} x_{1} x_{2}+\beta_{6} x_{1} x_{3}+$ $\beta_{7} x_{1} x_{4}$
(c) $H_{0}: \beta_{5}=\beta_{6}=\beta_{7}=0$ v.s. $H_{1}$ : at least one is not.

F -value $=\frac{(422.336-346.65) / 3}{346.63 / 32}=2.3297<F_{(3,32)}^{-1}(0.95)=2.9011$
We don't have enough evidence to state that for different location, type of packaging influence the total weekly sales.
11. Exercise 6.1
(a) F-Statistic is equivalent to T-Statistic for one variable case. So we compare the absolute value of T-Statistic $\left|\frac{\hat{\beta}}{S_{\beta}}\right|$
$\left|t\left(x_{1}\right)\right|=3.81,\left|t\left(x_{2}\right)\right|=90,\left|t\left(x_{3}\right)\right|=2.98$
$\left|t\left(x_{4}\right)\right|=1.21,\left|t\left(x_{5}\right)\right|=6.03,|t(x)|=0.86$
$x_{2}$ is the best candidate.

(b) Yes. Reject region is $|t(x)|>2.0106 . x_{2}$ should be included in model.
(c) (i) Fit all two- ariable model as: $y=\beta_{0}+\beta_{1} x_{2}+\beta_{2} x^{2}$ where $i=$ $1,3,4,5,6$.
(ii) Check all fitted two-variable model with a reduced model: $y=80+$ $\beta_{1} x / 2$ compare F -value.
(iii) keep the two-variable model with largest F-value, check all included variable.
12. Exercise 6.10
take the Traditional level as base level, represent the variables engine as twovalued variable.

$$
\begin{gathered}
\text { ADVANCED }= \begin{cases}1, & \text { at advanced level; } \\
0, & \text { not at advanced level. }\end{cases} \\
\text { AERODERIVATIVE }= \begin{cases}1, & \text { at aeroderivative level; } \\
0, & \text { not at aeroderivative level. }\end{cases}
\end{gathered}
$$

